

Geostatistical Modeling for Large Data Sets: Low-rank methods

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Outline

Motivation

Low-rank methods

Fixed Rank Kriging & Gaussian Predictive Process

Gaussian process (GP) geostatistics

Model:

$$Y(\mathbf{s}) = \mu(\mathbf{s}) + \eta(\mathbf{s}) + \varepsilon(\mathbf{s}), \quad \mathbf{s} \in \mathcal{S} \subset \mathbb{R}^d$$

where

$$\mu(\mathbf{s}) = \mathbb{E}[Y(\mathbf{s})] = \mathbf{X}(\mathbf{s})^T \boldsymbol{\beta}$$

$$\{\eta(\mathbf{s})\}_{\mathbf{s} \in \mathcal{S}} \sim GP(0, C(\cdot, \cdot))$$

$$\varepsilon(\mathbf{s}) \stackrel{\text{i.i.d.}}{\sim} N(0, \tau^2) \quad \forall \mathbf{s} \in \mathcal{S}$$

Likelihood

When the process $Y(\mathbf{s})$ is observed at n locations, say, $\mathbf{s}_1, \dots, \mathbf{s}_n$
i.e.,

$$\mathbf{Y} = (Y(\mathbf{s}_1), \dots, Y(\mathbf{s}_n))^T$$

Log-likelihood:

$$\ell_n \propto -\frac{1}{2} \log |\boldsymbol{\Sigma}_Y| - \frac{1}{2} (\mathbf{Y} - \mathbf{X}^T \boldsymbol{\beta})^T [\boldsymbol{\Sigma}_Y]^{-1} (\mathbf{Y} - \mathbf{X} \boldsymbol{\beta})$$

where

$$\begin{aligned} \boldsymbol{\Sigma}_Y &= [\text{Cov}\{Y(\mathbf{s}_i), Y(\mathbf{s}_j)\}]_{i,j=1,\dots,n} \\ &= [C(\mathbf{s}_i, \mathbf{s}_j) + \tau^2 \mathbb{1}_{\{\mathbf{s}_i = \mathbf{s}_j\}}]_{i,j=1,\dots,n} \end{aligned}$$

“Big n Problem” in geostatistics

- ▶ Modern environmental instruments have produced a wealth of space–time data $\Rightarrow n$ is big
- ▶ Evaluation of the likelihood function involves factorizing large covariance matrices that generally requires
 - ▶ $\mathcal{O}(n^3)$ operations
 - ▶ $\mathcal{O}(n^2)$ memory
- ▶ Modeling strategies are needed to deal with large spatial data set.
 - ▶ parameter estimation \Rightarrow MLE, Bayesian
 - ▶ spatial interpolation \Rightarrow Kriging
 - ▶ multivariate spatial data ($np \times np$), spatio-temporal data ($nt \times nt$)

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Low-rank approximation

Goal: want to approximate $\mathbf{Y} = \mathbf{X}^T \boldsymbol{\beta} + \boldsymbol{\eta} + \boldsymbol{\varepsilon}$ into a lower-dimensional structure.

Hierarchical Representation (assume zero mean)

$$\mathbf{Y} = \boldsymbol{\eta} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\varepsilon} \sim N_n(\mathbf{0}, \boldsymbol{\Sigma}_{\boldsymbol{\varepsilon}})$$

$$\boldsymbol{\eta} = \mathbf{AZ} + \boldsymbol{\xi}, \quad \boldsymbol{\xi} \sim N_n(\mathbf{0}, \boldsymbol{\Sigma}_{\boldsymbol{\xi}})$$

$$\mathbf{Z} \sim N_r(\mathbf{0}, \boldsymbol{\Sigma}_{\mathbf{Z}})$$

- ▶ $\mathbf{Z} = (Z_1, \dots, Z_r)^T$, $r \ll n$
- ▶ \mathbf{A} is the $(n \times r)$ expansion matrix that maps \mathbf{Z} to $\boldsymbol{\eta}$
- ▶ $\boldsymbol{\Sigma}_{\boldsymbol{\varepsilon}}$ and $\boldsymbol{\Sigma}_{\boldsymbol{\xi}}$ are (usually assumed to be) diagonal
- ▶ $\boldsymbol{\eta} = \mathbf{AZ} + \boldsymbol{\xi} \Rightarrow \eta(\mathbf{s}_i) = \sum_{j=1}^r a_{ij} Z_j + \xi(\mathbf{s}_i)$

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Computational savings using low-rank approximation

To carry out parameter estimation and spatial interpolation one need to compute

$$(A\Sigma_ZA^T + V)^{-1}$$

where $V = \Sigma_\epsilon + \Sigma_\xi$.

Sherman–Morrison–Woodbury formula

$$(A + BCD)^{-1} = A^{-1} - A^{-1}B(C^{-1} + DA^{-1}B)^{-1}DA^{-1}$$

In low-rank approximation, we have

$$(A\Sigma_ZA^T + V)^{-1} = V^{-1} - V^{-1}A(\Sigma_Z^{-1} + A^TV^{-1}A)^{-1}A^TV^{-1}$$

Low-rank approximation: expansion matrix A

$$\underbrace{\begin{bmatrix} \eta(\mathbf{s}_1) \\ \eta(\mathbf{s}_2) \\ \vdots \\ \vdots \\ \vdots \\ \eta(\mathbf{s}_n) \end{bmatrix}}_{\boldsymbol{\eta}} \approx \underbrace{\begin{bmatrix} a_{11} & a_{12} & \dots & a_{1r} \\ a_{21} & a_{22} & \dots & a_{2r} \\ \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots & \vdots \\ a_{n1} & a_{n2} & \dots & a_{nr} \end{bmatrix}}_{\mathbf{A}} \underbrace{\begin{bmatrix} Z_1 \\ Z_2 \\ \vdots \\ Z_r \end{bmatrix}}_{\mathbf{Z}}$$

Question

- ▶ How to choose A ?
- ▶ How to determine $\mathbf{Z} = (Z_1, \dots, Z_r)^T$?
 \Rightarrow choose r "knots" $\mathbf{s}_1^*, \dots, \mathbf{s}_r^* \in \mathcal{S}$ and let
 $\mathbf{Z} = (Z(\mathbf{s}_1^*), \dots, Z(\mathbf{s}_r^*))^T$

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Choice of A

- ▶ Orthogonal basis functions: e.g., Fourier, orthogonal polynomial, etc

- ▶ Karhune–Loève expansion:

$$\text{Cov}(\eta(\mathbf{s}_1), \eta(\mathbf{s}_2)) = \sum_{j=1}^{\infty} \lambda_j \phi_j(\mathbf{s}_1) \phi_j(\mathbf{s}_2) \text{ where}$$

$$\int_{\mathcal{S}} \text{Cov}(\eta(\mathbf{s}_1), \eta(\mathbf{s}_2)) \phi_j(\mathbf{s}_1) d\mathbf{s}_1 = \lambda_j \phi_j(\mathbf{s}_2)$$

- ▶ Empirical orthogonal functions (EOFs)
- ▶ Nonorthogonal Basis Functions: represent the spatial process as combination of kernel functions

$$\eta(\mathbf{s}) = \sum_{j=1}^r a(\mathbf{s}, \mathbf{s}_j^*) Z_j + \xi(\mathbf{s})$$

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Fixed Rank Kriging (Cressie & Johannesson 08)

Fixed-rank kriging outlines how kriging (i.e., spatial best linear unbiased predictor) can be applied in the low-rank parameterization of a spatial process.

- ▶ \mathbf{A} : (multiresolutional) non orthogonal basis functions

$$a(\mathbf{s})^T = \left(\{1 - (\|\mathbf{s} - \mathbf{s}_{1(l)}^*\|/\rho_{1(l)})^2\}_+^2, \dots, \{1 - (\|\mathbf{s} - \mathbf{s}_{r(l)}^*\|/\rho_{r(l)})^2\}_+^2 \right)$$

- ▶ $\mathbf{Z} = (Z(\mathbf{s}_1^*), \dots, Z(\mathbf{s}_r^*))^T$, $\Sigma_{\mathbf{Z}}$ to be estimated from data

The fixed rank kriging is equivalent to the following low rank model

$$Y(\mathbf{s}) = \mathbf{X}(\mathbf{s})^T \boldsymbol{\beta} + \sum_{j=1}^r a(\mathbf{s} - \mathbf{s}_{j(l)}^*) Z_j + \varepsilon(\mathbf{s}) + \xi(\mathbf{s})$$

Multiresolutional basis functions

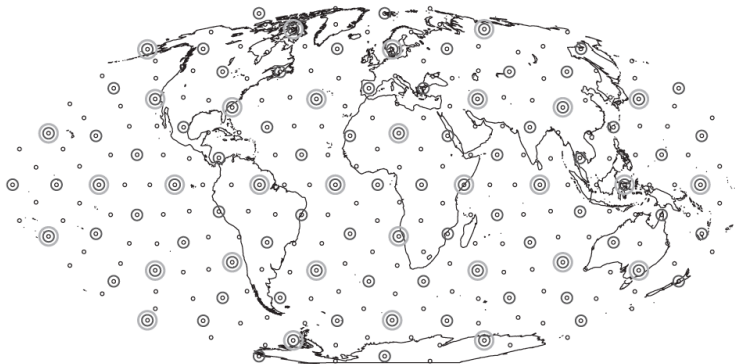


Figure: courtesy of Cressie & Johannesson 08

Nonstationary covariance function

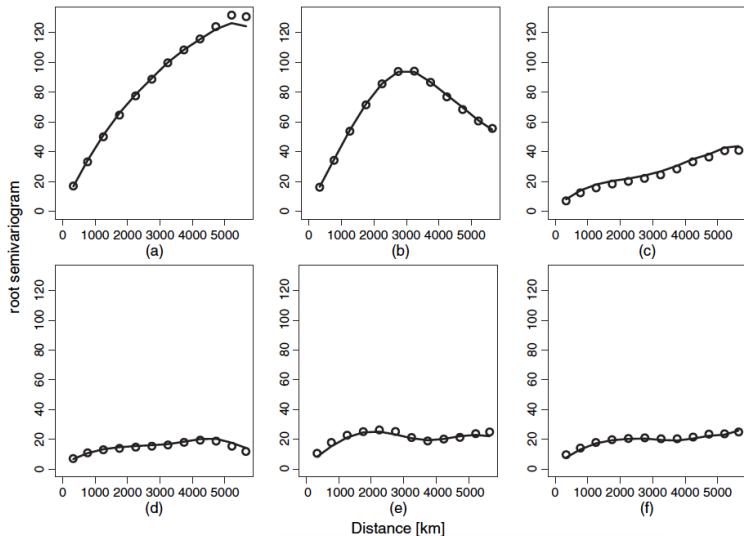


Figure: courtesy of Cressie & Johannesson 08

Gaussian Predictive Process (Banerjee et al. 08)

$$Y(\mathbf{s}) = \mathbf{X}(\mathbf{s})^T \boldsymbol{\beta} + \eta(\mathbf{s}) + \varepsilon(\mathbf{s}), \quad \eta \sim GP(0, C(\cdot, \cdot; \boldsymbol{\theta}))$$

Reasoning: Given \mathbf{Z} , we seek a linear combination $a(\mathbf{s})^T \mathbf{Z}$ such that

$$\int_{\mathcal{S}} \mathbb{E} \left[\eta(\mathbf{s}) - a(\mathbf{s})^T \mathbf{Z} \right]^2 d\mathbf{s}$$

is minimized.

In **predictive process** one let $\mathbf{Z} = \boldsymbol{\eta}^* = \{\eta(\mathbf{s}_i^*)\}_{i=1}^p \sim N_p(\mathbf{0}, C^*(\boldsymbol{\theta}))$

where $C^*(\boldsymbol{\theta}) = \left[C(\mathbf{s}_i^*, \mathbf{s}_j^*; \boldsymbol{\theta}) \right]_{i,j=1}^p$ and

$$\mathbf{A} = \left[C(\mathbf{s}_i, \mathbf{s}_j^*; \boldsymbol{\theta}) \right]_{i=1, \dots, n}^{j=1, \dots, p} [C^*]^{-1}$$

Covariance function approximation

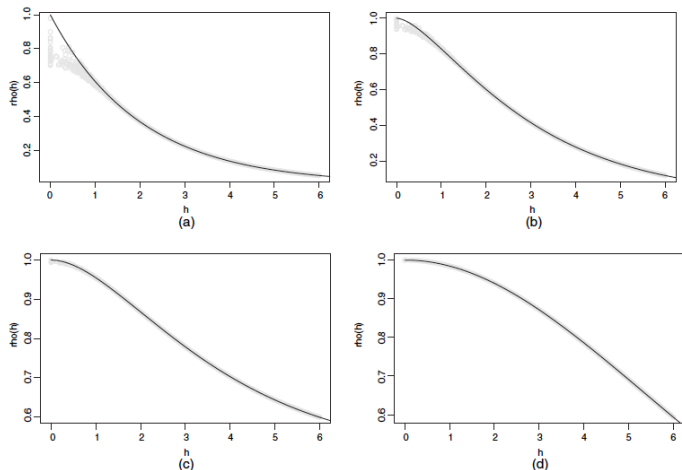


Fig. 1. Covariances of $w(s)$ against distance (—) and covariances of $\tilde{w}(s)$ against distance (\bullet): (a) smoothness parameter 0.5; (b) smoothness parameter 1; (c) smoothness parameter 1.5; (d) smoothness parameter 5

Figure: courtesy of Banerjee et al. 08

Covariance function approximation cont'd

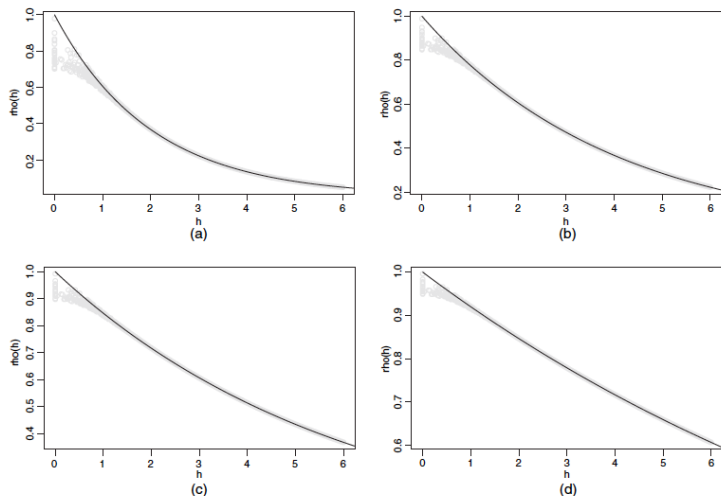





Fig. 2. Covariances of $w(\mathbf{s})$ against distance (—) and covariances of $\tilde{w}(\mathbf{s})$ against distance: (a) range parameter 2; (b) range parameter 4; (c) range parameter 6; (d) range parameter 12

Figure: courtesy of Banerjee et al. 08

Discussion

References

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