

Lecture 11

ARMA Models: Prediction and Forecasting

Reading: Bowerman, O'Connell, and Koehler (2005): Chapter 10.3; Cryer and Chen (2008): Chapter 9.1, 9.3, 9.4

MATH 4070: Regression and Time-Series Analysis

Linear Predictor

Prediction Equations

Examples

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Agenda

1 Linear Predictor

2 Prediction Equations

3 Examples

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Forecasting Stationary Time Series

Let $\{X_t\}$ be a **stationary process** with mean μ and ACVF $\gamma(\cdot)$. Based on the observed data, $\mathbf{X}_n = (X_1, X_2, \dots, X_n)^T$, we want to forecast X_{n+h} for some h , a positive integer

- **Question:** What is the best way to do so?
⇒ Need to decide on what “best” means

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- **Question:** What is the best way to do so?
⇒ Need to decide on what “best” means
- A commonly used metric for describing forecast performance is the **mean squared prediction error** (MSPE):

$$\text{MSPE} = \mathbb{E} \left[(X_{n+h} - m_n(\mathbf{X}_n))^2 \right].$$

⇒ the best predictor (in terms of MSPE) is

$$m_n(\mathbf{X}_n) = \mathbb{E} [X_{n+h} | \mathbf{X}_n],$$

the conditional expectation of X_{n+h} given \mathbf{X}_n

Linear Predictor

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Linear Predictor

Calculating $\mathbb{E}[X_{n+h} | \mathbf{X}_n]$ can be difficult in general

- We will restrict to a linear combination of X_1, X_2, \dots, X_n and a constant \Rightarrow **linear predictor**:

$$\begin{aligned} P_n X_{n+h} &= c_0 + c_1 X_n + c_2 X_{n-1} + \dots + c_n X_1 \\ &= c_0 + \sum_{j=1}^n c_j X_{n+1-j} \end{aligned}$$

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- We select the coefficients that minimize the **h -step-ahead mean squared prediction error**:

$$\mathbb{E}([X_{n+h} - P_n X_{n+h}]^2) = \mathbb{E}\left(X_{n+h} - c_0 - \sum_{j=1}^n c_j X_{n+1-j}\right)^2$$

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- The **best linear predictor** is the **best predictor** if $\{X_t\}$ is **Gaussian**

How to Determine these Coefficients $\{c_j\}$?

The steps that we are about to follow to calculate the c_j values are the same as you would use for calculating **ordinary least squares estimates**

- 1 Take the derivative of the MSPE with respect to each coefficient c_j

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- 2 Set each derivative equal to zero

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- 1 Take the derivative of the MSPE with respect to each coefficient c_j
- 2 Set each derivative equal to zero
- 3 Solve with respect to the coefficients

Linear Predictor

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Forecasting Stationary Processes I

For simplicity, let's assume $\mu = 0$ (we can always achieve that by subtracting off μ) so that we don't need the constant term. We have

$$P_n X_{n+h} = c_1 X_n + c_2 X_{n-1} + \cdots + c_n X_1.$$

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We want the MSPE

$$\mathbb{E}[(X_{n+h} - P_n X_{n+h})^2] = \mathbb{E}[(X_{n+h} - c_1 X_n - c_2 X_{n-1} - \cdots - c_n X_1)^2]$$

as small as possible.

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From now on let's definite

$$\mathbb{E}[(X_{n+h} - c_1 X_n - c_2 X_{n-1} - \dots - c_n X_1)^2] = S(c_1, \dots, c_n)$$

We are going to take derivative of the $S(c_1, \dots, c_n)$ with respect to each coefficient c_j

Linear Predictor

Prediction Equations

Examples

S is a quadratic function of c_1, c_2, \dots, c_n , so any minimizing set of c_j 's must satisfy these n equations:

$$\frac{\partial S(c_1, \dots, c_n)}{\partial c_j} = 0, \quad j = 1, \dots, n.$$

Recall $S(c_1, \dots, c_n) = \mathbb{E} [(X_{n+h} - c_1 X_n - c_2 X_{n-1} - \dots - c_n X_1)^2]$, we have

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$$\frac{\partial S(c_1, \dots, c_n)}{\partial c_j} = -2\mathbb{E} \left[\left(X_{n+h} - \sum_{i=1}^n c_i X_{n-i+1} \right) X_{n-j+1} \right] = 0$$

$$\Rightarrow \text{Cov}(X_{n+h} - \sum_{i=1}^n c_i X_{n-i+1}, X_{n-j+1}) = 0, \quad j = 1, \dots, n$$

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\Rightarrow Prediction error is uncorrelated with all RVs used in corresponding predictor

Linear Predictor

Prediction Equations

Examples

Orthogonality principle:

$$\text{Cov}\left(X_{n+h} - \sum_{i=1}^n c_i X_{n-i+1}, X_{n-j+1}\right) = 0, \quad j = 1, \dots, n.$$

Linear Predictor

Prediction Equations

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We have

$$\text{Cov}(X_{n+h}, X_{n-j+1}) - \sum_{i=1}^n c_i \text{Cov}(X_{n-i+1}, X_{n-j+1}) = 0$$

Linear Predictor

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We have

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We obtain $\{c_i; i = 1, \dots, n\}$ by solving the system of linear equations:

$$\left\{ \gamma(h+j-1) = \sum_{i=1}^n c_i \gamma(i-j) : j = 1, \dots, n \right\},$$

to find n unknown c_i 's

Linear Predictor

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We can rewrite the system of prediction equations as

$$\boldsymbol{\gamma}_n = \boldsymbol{\Sigma}_n \mathbf{c}_n,$$

with $\boldsymbol{\gamma}_n = (\gamma(h), \gamma(h+1), \dots, \gamma(h+n-1))^T$, $\mathbf{c}_n = (c_1, c_2, \dots, c_n)^T$
and

$$\boldsymbol{\Sigma}_n = \begin{bmatrix} \gamma(0) & \gamma(1) & \cdots & \gamma(n-1) \\ \gamma(1) & \gamma(0) & \cdots & \gamma(n-2) \\ \vdots & \vdots & \ddots & \vdots \\ \gamma(n-1) & \gamma(n-2) & \cdots & \gamma(0) \end{bmatrix}$$

is the covariance matrix of $(X_1, X_2, \dots, X_n)^T$.

Solving for \mathbf{c}_n we have

$$\mathbf{c}_n = \boldsymbol{\Sigma}_n^{-1} \boldsymbol{\gamma}_n$$

Linear Predictor

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Examples

The prediction errors are

$$\begin{aligned}U_{n+h} &= X_{n+h} - P_n X_{n+h} \\ &= (X_{n+h} - \mu) - \sum_{j=1}^n c_j (X_{n+1-j} - \mu).\end{aligned}$$

It then follows that

- The prediction error has mean zero

$$\mathbb{E}(U_{n+h}) = \mathbb{E}(X_{n+h} - P_n X_{n+h}) = 0$$

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It then follows that

- The prediction error has mean zero

$$\mathbb{E}(U_{n+h}) = \mathbb{E}(X_{n+h} - P_n X_{n+h}) = 0$$

- The prediction error is uncorrelated with all RVs used in the predictor

$$\text{Cov}(U_{n+h}, X_j) = \text{Cov}(X_{n+h} - P_n X_{n+h}, X_j) = 0, \quad j = 1, \dots, n$$

The Minimum Mean Squared Prediction Error

We obtain the minimum value of the MSPE by substituting the expression for \mathbf{c}_n into $\mathbb{E}[(X_{n+h} - P_n X_{n+h})^2]$:

$$\begin{aligned}
 \text{MSPE} &= \mathbb{E}[(X_{n+h} - P_n X_{n+h})^2] \\
 &= \mathbb{E}[(X_{n+h} - \mu)^2] - 2 \sum_{j=1}^n c_j \mathbb{E}[(X_{n+1-j} - \mu)(X_{n+h} - \mu)] \\
 &\quad + \mathbb{E}\left[\sum_{j=1}^n c_j (X_{n+1-j} - \mu)\right]^2 \\
 &= \mathbb{E}[(X_{n+h} - \mu)^2] - 2 \sum_{j=1}^n c_j \mathbb{E}[(X_{n+1-j} - \mu)(X_{n+h} - \mu)] \\
 &\quad + \sum_{j=1}^n \sum_{k=1}^n c_j c_k \mathbb{E}[(X_{n+1-j} - \mu)(X_{n+1-k} - \mu)] \\
 &= \gamma(0) - 2 \sum_{j=1}^n c_j \gamma(h+j-1) + \sum_{j=1}^n \sum_{k=1}^n c_j c_k \gamma(k-j) \\
 &= \gamma(0) - 2\mathbf{c}_n^T \boldsymbol{\gamma}_n + \mathbf{c}_n^T \boldsymbol{\Sigma}_n \mathbf{c}_n.
 \end{aligned}$$

From the previous slide we have

$$\text{MSPE} = \gamma(0) - 2\mathbf{c}_n^T \boldsymbol{\gamma}_n + \mathbf{c}_n^T \Sigma_n \mathbf{c}_n$$

Recall that $\mathbf{c}_n = \Sigma_n^{-1} \boldsymbol{\gamma}_n$, therefore we have

$$\begin{aligned} \text{MSPE} &= \gamma(0) - 2\mathbf{c}_n^T \boldsymbol{\gamma}_n + \mathbf{c}_n^T \Sigma_n \Sigma_n^{-1} \boldsymbol{\gamma}_n \\ &= \gamma(0) - \mathbf{c}_n^T \boldsymbol{\gamma}_n \\ &= \gamma(0) - \sum_{j=1}^n c_j \gamma(h+j-1). \end{aligned}$$

If $\{X_t\}$ is a Gaussian process then an **approximate** **100(1 - α)% prediction interval** for X_{n+h} is given by

$$P_n X_{n+h} \pm z_{1-\alpha/2} \sqrt{\text{MSPE}}.$$

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One-Step Ahead Prediction of AR(1) Process

Consider AR(1) process $X_t = \phi X_{t-1} + Z_t$, where $|\phi| < 1$ and $\{Z_t\} \sim \text{WN}(0, 1 - \phi^2)$.

- Since $\text{Var}(X_t) = 1$, $\gamma(h) = \rho(h) = \phi^{|h|}$
- To forecast X_{n+1} based upon $\mathbf{X}_n = (X_1, \dots, X_n)^T$, using best linear predictor $P_n X_{n+1} = \mathbf{c}_n^T \mathbf{X}_n$, we need to solve $\Sigma_n \mathbf{c}_n = \gamma_n$

$$\begin{bmatrix} 1 & \phi & \dots & \phi^{n-1} \\ \phi & 1 & \dots & \phi^{n-2} \\ \vdots & \vdots & \dots & \vdots \\ \phi^{n-1} & \phi^{n-2} & \dots & 1 \end{bmatrix} \begin{bmatrix} c_1 \\ c_2 \\ \vdots \\ c_n \end{bmatrix} = \begin{bmatrix} \phi \\ \phi^2 \\ \vdots \\ \phi^n \end{bmatrix}$$

\Rightarrow the solution is $\mathbf{c}_n = (\phi, 0, \dots, 0)^T$, yielding

$$P_n X_{n+1} = \mathbf{c}_n^T \mathbf{X}_n = \phi X_n$$

- ϕX_n makes intuitive sense as a predictor since

$$X_{n+1} = \phi X_n + Z_{n+1}$$

Linear Predictor

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- ϕX_n makes intuitive sense as a predictor since

$$X_{n+1} = \phi X_n + Z_{n+1}$$

- Prediction error is $X_{n+1} - \phi X_n = Z_{n+1}$ and

$$\text{Cov}(Z_t, X_{n-j+1}) = 0, j = 1, \dots, n$$

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$$\text{Cov}(Z_t, X_{n-j+1}) = 0, j = 1, \dots, n$$

- MSPE is

$$\text{Var}(X_{n+1} - \phi X_n) = \gamma(0) - \mathbf{c}_n^T \boldsymbol{\gamma}_n = 1 - \phi^2,$$

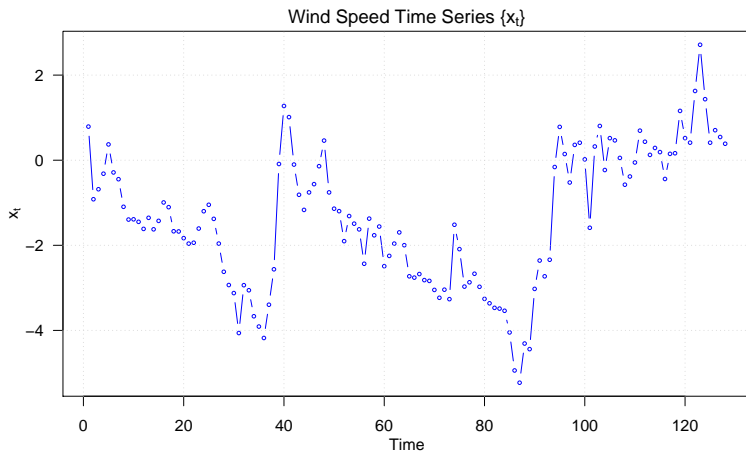
because $\mathbf{c}_n = (\phi, 0, \dots, 0)^T$ and $\boldsymbol{\gamma}_n = (\phi, \phi^2, \dots, \phi^n)^T$

Linear Predictor

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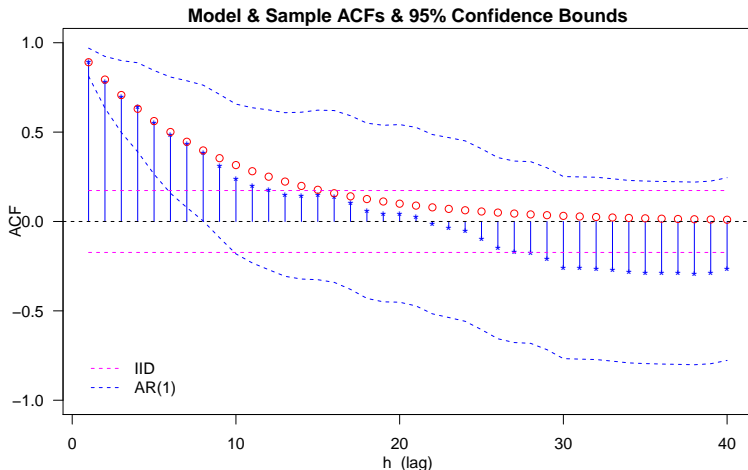
Examples

Wind Speed Time Series Example [Source: UW stat 519 lecture notes by Donald Percival]



Let's use this series to illustrate forecasting one step ahead

Model & Sample ACFs & 95% Confidence Bounds



The sample ACF indicates compatibility with AR(1) model

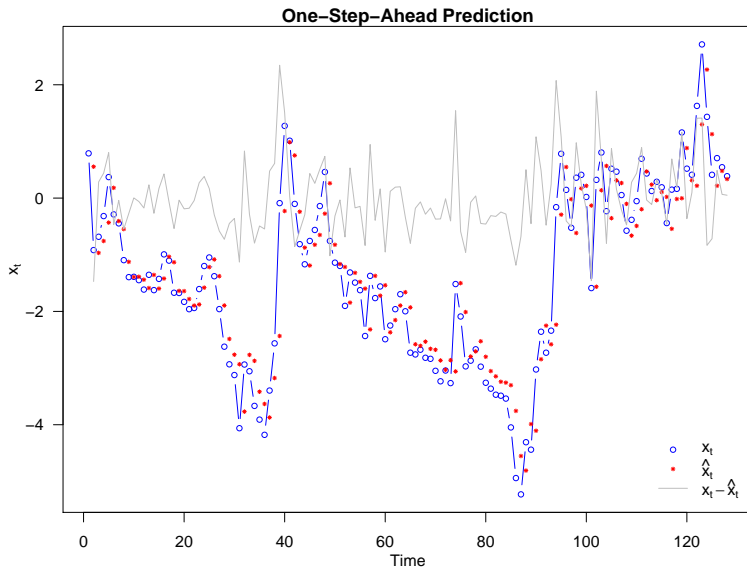
$$\Rightarrow P_n X_{n+1} = \phi X_n$$

One-Step-Ahead Prediction of Wind Speed Series

Linear Predictor

Prediction Equations

Examples



Predicting “Missing” Values

- Let $\{X_t\}$ be a stationary process with mean μ and ACVF $\gamma(\cdot)$. Suppose we know X_1 and X_3 , and want to predict X_2 using linear combinations of X_1 and X_3



Linear Predictor

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- Let $\{X_t\}$ be a stationary process with mean μ and ACVF $\gamma(\cdot)$. Suppose we know X_1 and X_3 , and want to predict X_2 using linear combinations of X_1 and X_3
- Solution:** To calculate $P_{X_1, X_3} X_2$ we minimize

$$\begin{aligned} \text{MSPE} &= \mathbb{E} \left[(X_2 - P_{X_1, X_3} X_2)^2 \right] \\ &= \mathbb{E} \left[(X_2 - c_0 - c_1 X_3 - c_2 X_1)^2 \right] \end{aligned}$$

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- Proceed as for the forecasting case to get the optimal coefficients:

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 - Calculate derivatives

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 - Calculate derivatives
 - Set the derivatives equal to zero

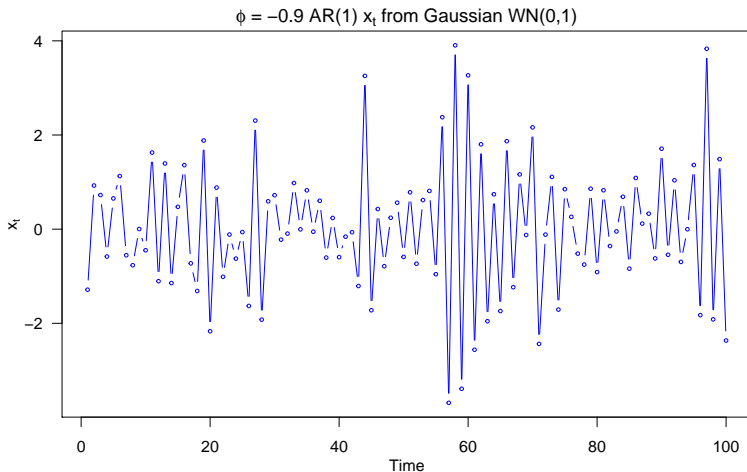
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- Proceed as for the forecasting case to get the optimal coefficients:
 - Calculate derivatives
 - Set the derivatives equal to zero
 - Solve the linear system of equation

Another AR(1) Example with $\phi = -0.9$

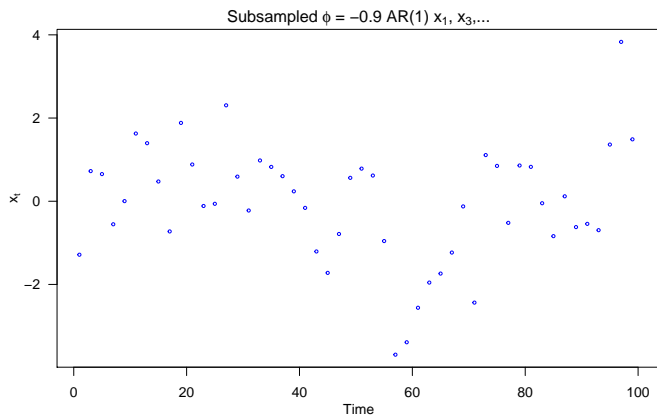


Linear Predictor

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Subsampled X_1, X_3, \dots and Removed X_2, X_4, \dots



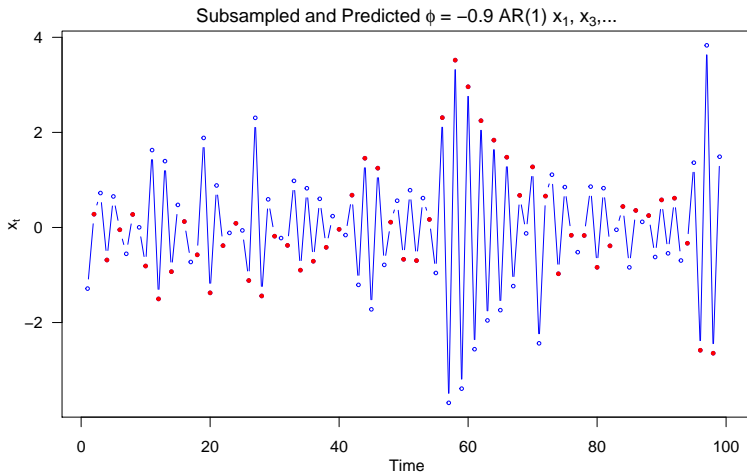
The best linear predictor of X_2 given X_1, X_3 is

$$\hat{X}_2 = \frac{\phi}{1 + \phi^2} (X_1 + X_3),$$

and the MSPE is

$$\frac{\sigma^2}{1 + \phi^2}$$

Predict X_2, X_4, \dots Using Best Linear Predictor

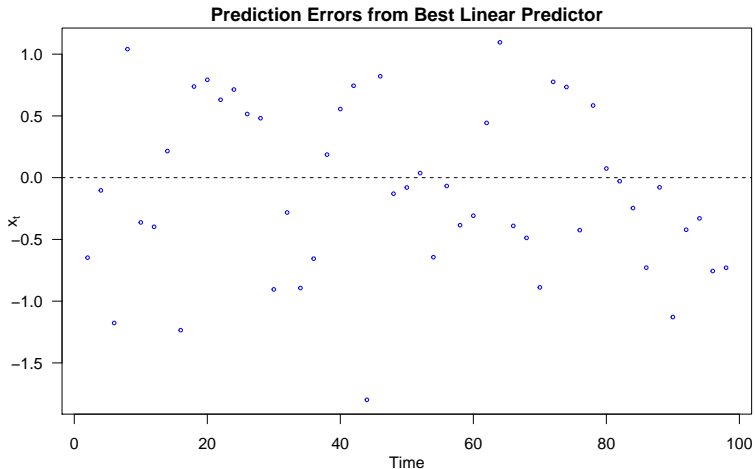


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Prediction Errors from Best Linear Predictor



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