

Lecture 6

Autocorrelation and Time Series Models

Reading: Forecasting, Time Series, and Regression (4th
edition) by Bowerman, O'Connell, and Koehler: Chapter 6

MATH 4070: Regression and Time-Series Analysis

Objectives of Time
Series Analysis

Time Series Models

Mean and
Autocovariance
Functions

Stationarity

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2 Time Series Models

3 Mean and Autocovariance Functions

4 Stationarity

Objectives of Time
Series Analysis

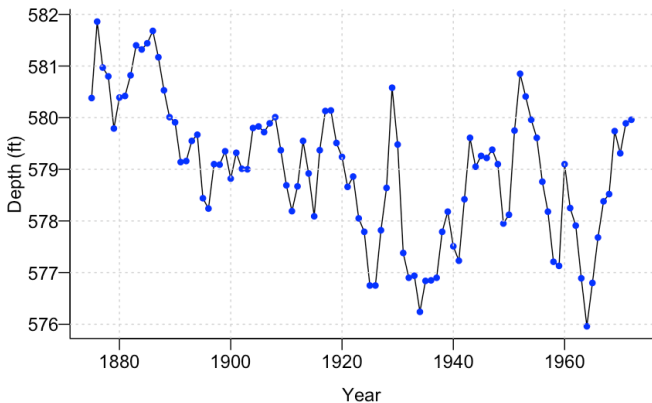
Time Series Models

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Level of Lake Huron 1875–1972

```
```{r}  
par(mar = c(3.2, 3.2, 0.5, 0.5), mgp = c(2, 0.5, 0), bty = "L")
data(LakeHuron)
plot(LakeHuron, ylab = "Depth (ft)", xlab = "Year", las = 1)
points(LakeHuron, cex = 0.8, col = "blue", pch = 16)
grid()
```
```



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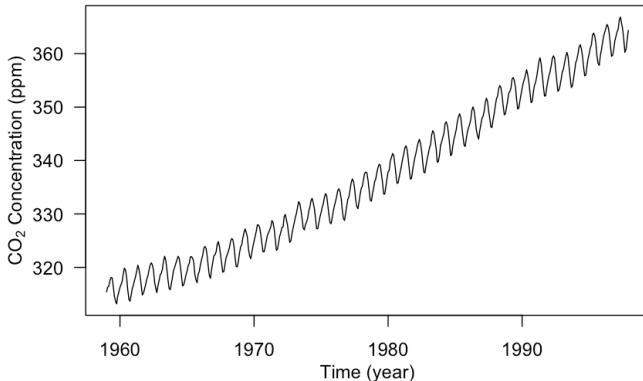
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Mauna Loa Atmospheric CO₂ Concentration

```
```{r}  
data(co2)
par(mar = c(3.8, 4, 0.8, 0.6))
plot(co2, las = 1, xlab = "", ylab = "")
mtext("Time (year)", side = 1, line = 2)
mtext(expression(paste("CO"[2], " Concentration (ppm)")), side = 2, line = 2.5)
```
```



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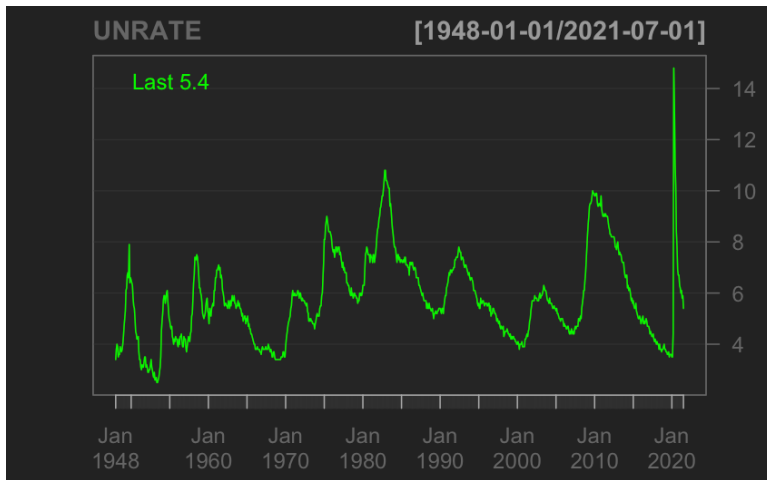
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US Unemployment Rate 1948 Jan. – 2021 July

[Source: St. Louis Federal Reserve Bank's FRED system]



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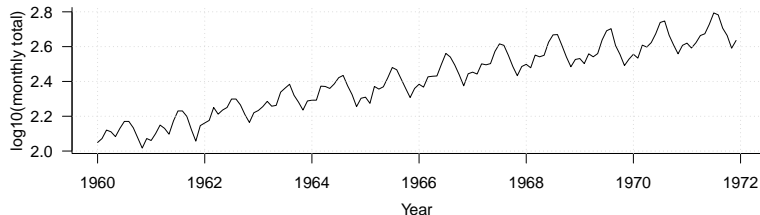
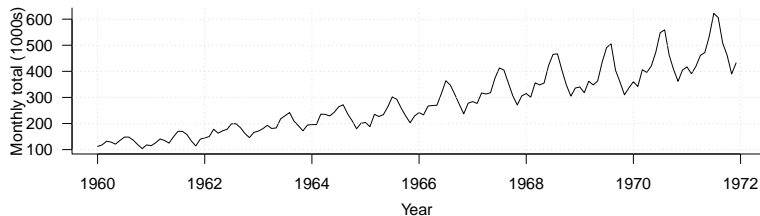
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Airline Passengers Example

The data set `airpassengers`, which are the monthly totals of international airline passengers from 1960 to 1971.



Here we stabilize the variance with a \log_{10} transformation

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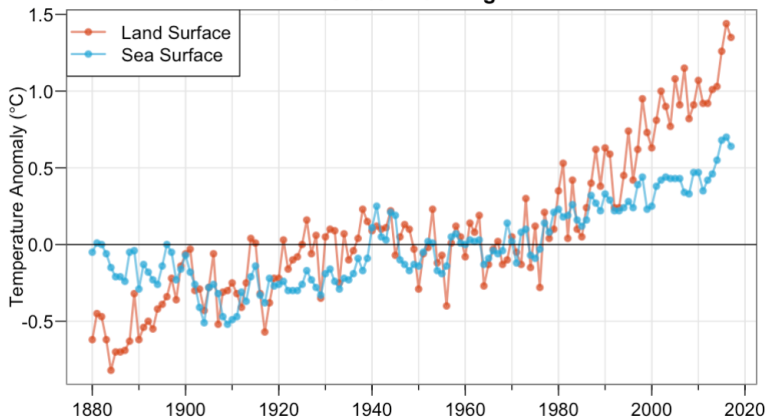
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Global Annual Temperature Anomalies

[Source: NASA GISS Surface Temperature Analysis]

Global Warming



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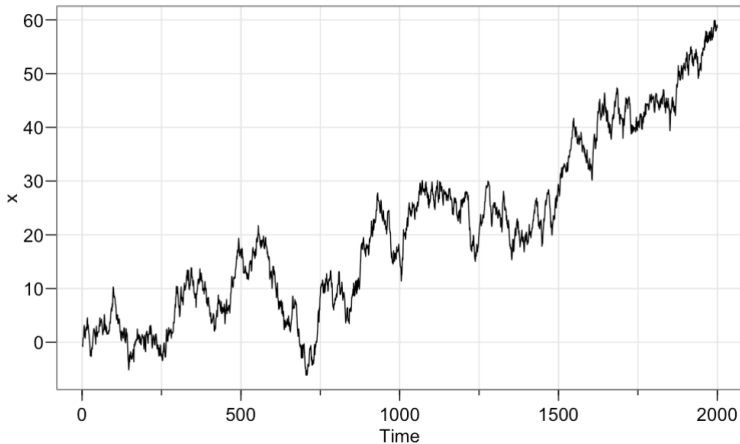
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A Simulated Time Series

```
```{r}  
set.seed(123)
w <- rnorm(2000); x <- cumsum(w); tsplot(x, las = 1)
```
```



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Statistical Modeling: Find a **statistical model** that adequately explains the observed time series

- For example, identify a model which can account for the fact that the depths of Lake Huron are correlated with different years and with a decreasing long-term trend

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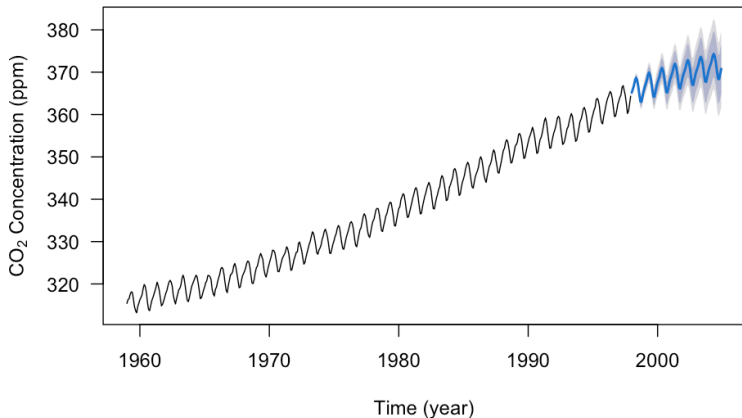
Statistical Modeling: Find a **statistical model** that adequately explains the observed time series

- For example, identify a model which can account for the fact that the depths of Lake Huron are correlated with different years and with a decreasing long-term trend
- The fitted model can be used for further **statistical inference**, for instance, to answer the question like: **Is there evidence of decreasing trend in the Lake Huron depths?**

Some Objectives of Time Series Analysis, Cont'd

Forecasting is perhaps the most common objective. One observe a time series of given length and wish to **predict** or **forecast** future values of the time series based on those already observed.

Forecasts from TBATS(1, {3,1}, -, {<12,5>})



- **Adjustment:** an example would be **seasonal adjustment**, where the seasonal component is estimated and then removed to better understand the underlying trend
- **Simulation:** use a time series model (which adequately describes a physical process) as a surrogate to *simulate repeatedly in order to approximate how the physical process behaves*
- **Control:** adjust various **input (control)** parameters to make the time series fit more closely to a given standard (many examples from statistical quality control)

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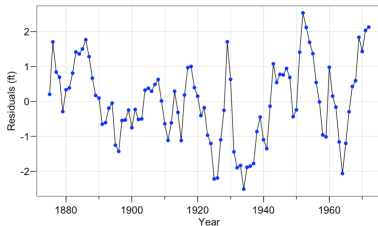
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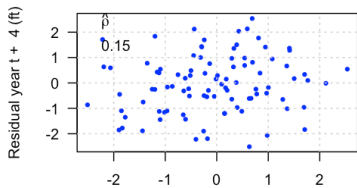
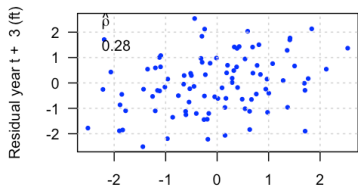
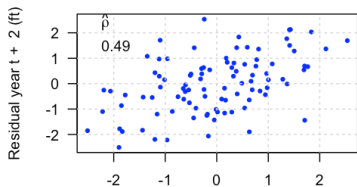
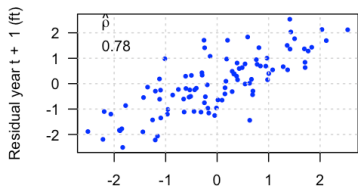
Lake Huron Time Series

- **Time series analysis** is the area of statistics which deals with the analysis of **dependency** between different observations (typically $\{\eta_t\}$)
- Some key features of the Lake Huron time series: 
 - decreasing trend
 - some “random” fluctuations around the decreasing trend
- For example, we can extract the ‘noise’ component by assuming a linear trend



Exploring the Dependence Structure of “Noise” $\{\eta_t\}$

$\{\eta_t\}$ exhibit a temporal dependence structure, meaning that the nearby (in time) values tend to be more alike than those that are far part. To observe this, let's create a few time lag plots

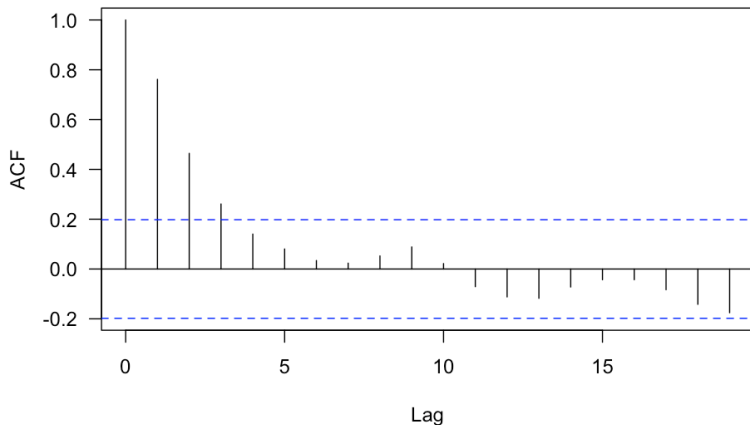


Residual year t (ft)

Residual year t (ft)

Further Exploration of the Temporal Dependence Structure

Let's plot the correlation as a function of the time lag



We will learn how to use this information to suggest an appropriate **time series model**

Time Series Models

- A **time series model** is a probabilistic model that describes how the series data y_t could have been generated. More specifically, it is a probability model for $\{Y_t : t \in T\}$, a **collection of random variables indexed in time**

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- We will keep our models for Y_t as simple as possible by assuming **stationarity**, meaning that some characteristics of the distribution of Y_t depend only on the “time lag” not on the specific time points
- While most time series are not stationary, we can model the non-stationary parts (e.g., by **de-trending** or **de-seasonalizing**) to obtain a stationary component, η_t . We typically assume the process is **second-order stationary**, meaning

$$\begin{aligned} \mathbb{E}[\eta_t] &= 0, \quad \forall t \in T \quad \text{and,} \\ \text{Cov}(\eta_t, \eta_{t'}) &= \gamma(t' - t) = \text{Cov}(\eta_{t+s}, \eta_{t'+s}) \end{aligned}$$



- A **time series model** is a specification of the probabilistic distribution of a sequence of random variables (RVs) η_t

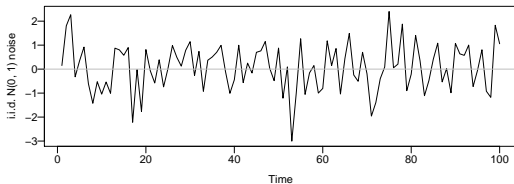
(The observed time series is a **realization** of such a sequence of random variables)

- The simplest time series is **i.i.d. (*independent and identically distributed*) noise**
 - $\{\eta_t\}$ is a sequence of independent and identically distributed zero-mean (i.e., $\mathbb{E}(\eta_t) = 0, \forall t$) random variables
 \Rightarrow **no temporal dependence**
 - It is of little value of using i.i.d. noise model to conduct **forecast** as there is no information from the past observations
 - **But**, we will use i.i.d. model as a building block to develop time series models that can accommodate time dependence

Example Realizations of i.i.d. Noise

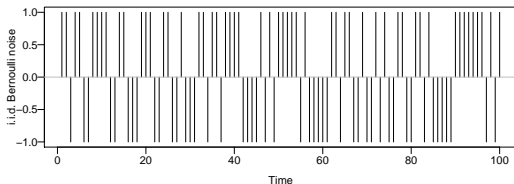
- Gaussian (normal) i.i.d. noise with mean 0 and variance $\sigma^2 > 0$

$$f(\eta_t) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{\eta_t^2}{2\sigma^2}\right)$$



- Bernoulli i.i.d. noise with “success” probability

$$\mathbb{P}(\eta_t = 1) = p = 1 - \mathbb{P}(\eta_t = -1)$$



A time series model could also be a specification of the **means** and **autocovariances** of the RVs

- The **mean function** of $\{\eta_t\}$ is

$$\mu_t = \mathbb{E}(\eta_t).$$

- μ_t is the population mean at time t , which can be computed as:

$$\mu_t = \begin{cases} \int_{-\infty}^{\infty} \eta_t f(\eta_t) d\eta_t & \text{when } \eta_t \text{ is a continuous RV;} \\ \sum_{-\infty}^{\infty} \eta_t p(\eta_t), & \text{when } \eta_t \text{ is a discrete RV,} \end{cases}$$

where $f(\cdot)$ and $p(\cdot)$ are the probability density function and probability mass function of η_t , respectively

- **Example 1:** What is the mean function for $\{\eta_t\}$, an i.i.d. $N(0, \sigma^2)$ process?

- **Example 2:** For each time point, let $Y_t = \beta_0 + \beta_1 t + \eta_t$ with β_0 and β_1 some constants and η_t is defined above. What is $\mu_Y(t)$?

Review: The Covariance Between Two RVs

- The **covariance** between the RVs X and Y is

$$\begin{aligned}\text{Cov}(X, Y) &= \mathbb{E}\{(X - \mu_X)(Y - \mu_Y)\} \\ &= \mathbb{E}(XY) - \mu_X \mu_Y.\end{aligned}$$

It is a measure of **linear dependence** between the two RVs. When $X = Y$ we have

$$\text{Cov}(X, X) = \text{Var}(X).$$

- For constants a, b, c , and RVs X, Y, Z :

$$\begin{aligned}\text{Cov}(aX + bY + c, Z) &= \text{Cov}(aX, Z) + \text{Cov}(bY, Z) \\ &= a\text{Cov}(X, Z) + b\text{Cov}(Y, Z)\end{aligned}$$

\Rightarrow

$$\begin{aligned}\text{Var}(X + Y) &= \text{Cov}(X, X) + \text{Cov}(X, Y) + \text{Cov}(Y, X) + \text{Cov}(Y, Y) \\ &= \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)\end{aligned}$$

- The autocovariance function of $\{\eta_t\}$ is

$$\gamma(s, t) = \text{Cov}(\eta_s, \eta_t) = \mathbb{E}[(\eta_s - \mu_s)(\eta_t - \mu_t)]$$

It measures the strength of linear dependence between two RVs η_s and η_t

- **Properties:**

- $\gamma(s, t) = \gamma(t, s)$ for each s and t
- When $s = t$ we have

$$\gamma(t, t) = \text{Cov}(\eta_t, \eta_t) = \text{Cov}(\eta_t) = \sigma_t^2$$

the value of the variance function at time t

- $\gamma(s, t)$ is a non-negative definite function (will come back to this later)

- The autocorrelation function of $\{\eta_t\}$ is

$$\rho(s, t) = \text{Corr}(\eta_s, \eta_t) = \frac{\gamma(s, t)}{\sqrt{\gamma(s, s)\gamma(t, t)}}$$

It measures the “scale invariant” linear association between η_s and η_t

- **Properties:**

- $-1 \leq \rho(s, t) \leq 1$ for each s and t
- $\rho(s, t) = \rho(t, s)$ for each s and t
- $\rho(t, t) = 1$ for each t
- $\rho(\cdot, \cdot)$ is a non-negative definite function

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- We typically need “replicates” to estimate population quantities. For example, we use

$$\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$$

to be the estimate of μ_X , the population mean of the **single** RV, X

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- However, in time series analysis, we have $n = 1$ (i.e., no replication) because we only have one realized value at each time point

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- However, in time series analysis, we have $n = 1$ (i.e., no replication) because we only have one realized value at each time point
- **Stationarity** means that some characteristic of $\{\eta_t\}$ does not depend on the time point, t , only on the “time lag” between time points **so that we can create “replicates”**

Next, we will discuss **strict stationarity** and **weak stationarity**

- A time series, $\{\eta_t\}$, is **strictly stationary** if

$$[\eta_1, \eta_2, \dots, \eta_T] \stackrel{d}{=} [\eta_{1+h}, \eta_{2+h}, \dots, \eta_{T+h}],$$

for all integers h and $T \geq 1 \Rightarrow$ the **joint distribution** are unaffected by time shifts

- Under such the strict stationarity
 - $\{\eta_t\}$ is **identically distributed** but not (necessarily) **independent**
 - $\mu_t = \mu$ is independent of time t
 - $\gamma(s, t) = \gamma(s + h, t + h)$, for any s, t , and h

- $\{\eta_t\}$ is weakly stationary if
 - $\mathbb{E}(\eta_t) = \mu_t = \mu$
 - $\text{Cov}(\eta_t, \eta_{t+h}) = \gamma(t, t+h) = \gamma(h)$, finite constant that can depend on h but not on t
- Other names for this type of stationarity include second-order, covariance, wide sense. The quantity h is called the lag
- Weak and strict stationarity
 - A strictly stationary process $\{\eta_t\}$ is also weakly stationary as long as μ is finite
 - Weak stationarity does not imply strict stationarity!

The autocovariance function (ACVF) of a stationary process $\{\eta_t\}$ is defined to be

$$\begin{aligned}\gamma(h) &= \text{Cov}(\eta_t, \eta_{t+h}) \\ &= \mathbb{E}[(\eta_t - \mu)(\eta_{t+h} - \mu)],\end{aligned}$$

which measures the lag- h time dependence

Properties of the ACVF:

- $\gamma(0) = \text{Var}(\eta_t)$
- $\gamma(-h) = \gamma(h)$ for each h
- $\gamma(s-t)$ as a function of $(s-t)$ is **non-negative definite**

The autocorrelation function (ACF) of a stationary process $\{\eta_t\}$ is defined to be

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)}$$

which measures the “scale invariant” lag- h time dependence

Properties of the ACF:

- $-1 \leq \rho(h) \leq 1$ and $\rho(0) = 1$ for each h
- $\rho(-h) = \rho(h)$ for each h
- $\rho(s - t)$ as a function of $(s - t)$ is non-negative definite

In this lecture, we discuss

- Objectives of time series analysis
- Time series models
- Mean and auto-covariance/correlation functions
- Stationarity assumption in time series

The most important \mathbb{R} function of this lecture is acf , which calculates and plots the **sample autocorrelation**

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