## Lecture 5

Multiple Linear Regression I
Reading: Chapter 12

## STAT 8020 Statistical Methods II

September 3, 2020

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Multiple Linear Regression (MLR)
Goal: To model the relationship between two or more explanatory variables ( $X$ 's) and a response variable ( $Y$ ) by fitting a linear equation to observed data:
$Y_{i}=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\cdots+\beta_{p-1} X_{p-1}+\varepsilon_{i}, \quad \varepsilon_{i} \stackrel{i . i . d .}{\sim} \mathrm{N}\left(0, \sigma^{2}\right)$

Example: Species diversity on the Galapagos Islands We are interested in studying the relationship between the number of plant species (Species) and the following geographic variables: Area, Elevation, Nearest, Scruz, Adjacent.


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Data: Species Diversity on the Galapagos Islands


How Do Geographic Variables Affect Species Diversity?


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Let's Take a Look at the Correlation Matrix


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Model 2: Species ~Elevation + Area


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Model 3: Species $\sim$ Elevation + Area + Adjacent


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"Full Model"


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MLR Topics
Similar to SLR, we will discuss

- Estimation
- Inference
- Diagnostics and Remedies

We will also discuss some new topics

- Model Selection

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Multiple Linear Regression

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Multiple Linear Regression I

Multiple Linear Regression in Matrix Notation Multiple Linear Regression (MLR):
$\left(\begin{array}{c}Y_{1} \\ Y_{2} \\ \vdots \\ Y_{n}\end{array}\right)=\left(\begin{array}{ccccc}1 & X_{1,1} & X_{2,1} & \cdots & X_{p-1,1} \\ 1 & X_{1,2} & X_{2,2} & \cdots & X_{p-1,2} \\ \vdots & \cdots & \ddots & \vdots & \vdots \\ 1 & X_{1, n} & X_{2, n} & \cdots & X_{p-1, n}\end{array}\right)\left(\begin{array}{c}\beta_{0} \\ \beta_{1} \\ \vdots \\ \beta_{p-1}\end{array}\right)+\left(\begin{array}{c}\varepsilon_{1} \\ \varepsilon_{2} \\ \vdots \\ \varepsilon_{n}\end{array}\right)$
We can express MLR as

$$
\boldsymbol{Y}=\boldsymbol{X} \boldsymbol{\beta}+\boldsymbol{\varepsilon}
$$

Error Sum of Squares (SSE)
$=\sum_{i=1}^{n}\left(Y_{i}-\beta_{0}-\sum_{j=1}^{p-1} \beta_{j} X_{j}\right)^{2}$ can be expressed in matrix notation as:

$$
(\boldsymbol{Y}-\boldsymbol{X} \boldsymbol{\beta})^{T}(\boldsymbol{Y}-\boldsymbol{X} \boldsymbol{\beta})
$$

Again, we are going to find $\hat{\beta}$ to minimize SSE as our estimate for $\boldsymbol{\beta}$

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Estimation of Regression Coefficients

- The resulting least squares estimate is

$$
\hat{\boldsymbol{\beta}}=\left(\boldsymbol{X}^{T} \boldsymbol{X}\right)^{-1} \boldsymbol{X}^{T} \boldsymbol{Y}
$$

- Fitted values:

$$
\hat{\boldsymbol{Y}}=\boldsymbol{X} \hat{\boldsymbol{\beta}}=\boldsymbol{X}\left(\boldsymbol{X}^{T} \boldsymbol{X}\right)^{-1} \boldsymbol{X}^{T} \boldsymbol{Y}=\boldsymbol{H} \boldsymbol{Y}
$$

- Residuals:

$$
e=\boldsymbol{Y}-\hat{\boldsymbol{Y}}=(\boldsymbol{I}-\boldsymbol{H}) \boldsymbol{Y}
$$

## Estimation of $\sigma^{2}$

- Similar approach as we did in SLR

$$
\begin{aligned}
\hat{\sigma}^{2} & =\frac{\boldsymbol{e}^{T} \boldsymbol{e}}{n-p} \\
& =\frac{(\boldsymbol{Y}-\boldsymbol{X} \hat{\boldsymbol{\beta}})^{T}(\boldsymbol{Y}-\boldsymbol{X} \hat{\boldsymbol{\beta}})}{n-p} \\
& =\frac{\text { SSE }}{n-p} \\
& =\text { MSE }
\end{aligned}
$$

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Testing Individual Predictor

- We can show that $\hat{\boldsymbol{\beta}} \sim \mathrm{N}_{p}\left(\boldsymbol{\beta}, \sigma^{2}\left(\boldsymbol{X}^{T} \boldsymbol{X}\right)^{-1}\right) \Rightarrow$ $\hat{\beta}_{k} \sim \mathrm{~N}\left(\beta_{k}, \sigma_{\hat{\beta}_{k}}^{2}\right)$


## - Perform t test:

- $H_{0}: \beta_{k}=0$ vs. $H_{a}: \beta_{k} \neq 0$
- $\frac{\hat{\beta}_{k}-\beta_{k}}{\hat{\sigma}_{\hat{\beta}_{k}}} \sim t_{n-p} \Rightarrow t^{*}=\frac{\hat{\beta}_{k}}{\hat{\sigma}_{\hat{\beta}_{k}}} \stackrel{H_{0}}{\sim} t_{n-p}$
- Reject $H_{0}$ if $\left|t^{*}\right|>t_{1-\alpha / 2, n-p}$
- Confidence interval for $\beta_{k}: \hat{\beta}_{k} \pm t_{1-\alpha / 2, n-p} \hat{\sigma}_{\hat{\beta}_{k}}$


## Coefficient of Determination

- Coefficient of Determination $R^{2}$ describes proportional of the variance in the response variable that is predictable from the predictors

$$
R^{2}=\frac{\mathrm{SSR}}{\mathrm{SST}}=1-\frac{\mathrm{SSE}}{\mathrm{SSR}}, \quad 0 \leq R^{2} \leq 1
$$

- $R^{2}$ usually increase with the increasing $p$, the number of the predictors
- Adjusted $R^{2}$, denoted by $R_{\text {adj }}^{2}=\frac{\mathrm{SSR} /(n-p)}{\mathrm{SST} /(n-1)}$ attempts to account for $p$
$R^{2}$ vs. $R_{\text {adj }}^{2}$ Example

Suppose the true relationship between response $Y$ and predictors $\left(X_{1}, X_{2}\right)$ is

$$
Y=5+2 X_{1}+\varepsilon
$$

where $\varepsilon \sim \mathrm{N}(0,1)$ and $X_{1}$ and $X_{2}$ are independent to each other. Let's fit the following two models to the "data"

Model 1: $Y=\beta_{0}+\beta_{1} X_{1}+\varepsilon^{1}$
Model 2: $Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\varepsilon^{2}$

Question: Which model will "win" in terms of $R^{2}$ ?

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Model 1 Fit
> summary(fit1)
Call:
$\operatorname{lm}($ formula $=\mathrm{y} \sim \mathrm{x} 1$ )
Residuals:
Min $\quad 10$ Median $\quad 30 \quad$ Max
$-1.6085-0.5056-0.2152 \quad 0.6932 \quad 2.0118$
Coefficients:

|  | Estimate | Error | value | 1) |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| (Intercept) | 5.1720 | 0.1534 | 33.71 | < 2e-16 | *** |
| $\times 1$ | 1.8660 | 0.1589 | 11.74 | $2.47 \mathrm{e}-12$ |  |
| Signif. code |  |  |  |  |  |
| '***’ 0.00 | $1^{\text {'**' }} 0$. | ** 0.05 | , 0.1 | ، , |  |

Residual standard error: 0.8393 on 28 degrees of freedom
Multiple R-squared: 0.8313, Adjusted R-squared: 0.8253 F-statistic: 138 on 1 and 28 DF , p -value: $2.467 \mathrm{e}-12$

## Model 2 Fit

> summary(fit2)
Call:
$\operatorname{lm}($ formula $=y \sim x 1+x 2)$

| Residuals: |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Min | $1 Q$ | Median | $3 Q$ | Max |
| -1.3926 | -0.5775 | -0.1383 | 0.5229 | 1.8385 |

Coefficients:

|  | Estimate Std. Error $t$ value $\operatorname{Pr}(>\|t\|)$ |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: |
| (Intercept) | 5.1792 | 0.1518 | 34.109 | $<2 \mathrm{e}-16^{* * *}$ |
| $\times 1$ | 1.8994 | 0.1593 | 11.923 | $2.88 \mathrm{e}-12^{* * *}$ |
| $\times 2$ | -0.2289 | 0.1797 | -1.274 | 0.213 |

Signif. codes:
$0^{\prime * * *} 0.001$ ‘**' 0.01 '*' 0.05 '.' 0.1 ', 1
Residual standard error: 0.8301 on 27 degrees of freedom
Multiple R-squared: 0.8408, Adjusted R-squared: 0.8291 F-statistic: 71.32 on 2 and 27 DF, p-value: $1.677 \mathrm{e}-11$


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Species Diversity on the Galapagos Islands Revisited: Full Model
> summary(gala_fit2)
Call:
m(formula = Species ~ Elevation + Area)

| Residuals: |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: |
| Min | $1 Q$ | Median | $3 Q$ | Max |
| -192.619 | -33.534 | -19.199 | 7.541 | 261.514 |

Coefficients
Estimate Std. Error $t$ value $\operatorname{Pr}(>|t|)$
Intercept) $17.10519 \quad 20.94211 \quad 0.817 \quad 0.42120$
$\begin{array}{lllll}\text { Elevation } & 0.17174 & 0.05317 & 3.230 & 0.00325 \text { ** }\end{array}$
$\begin{array}{lllll}\text { Area } & 0.01880 & 0.02594 & 0.725 & 0.47478\end{array}$
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' $0.05^{\prime}$ ' 0.1 ',
Residual standard error: 79.34 on 27 degrees of freedom Multiple R-squared: 0.554, Adjusted R-squared: 0.52 F-statistic: 16.77 on 2 and 27 DF, p-value: 1.843e-05

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Perform a General Linear Test

- $H_{0}: \beta_{\text {Area }}=0$ vs. $H_{a}: \beta_{\text {Area }} \neq 0$
- $F^{*}=\frac{(173254-169947) /(2-1)}{169947 /(30-2-1)}=0.5254$
- P-value: $\mathrm{P}[F>0.5254]=0.4748$, where $F \sim \mathrm{~F}(1,27)$
> anova(gala_fit1, gala_fit2)
Analysis of Variance Table

Model 1: Species ~ Elevation
Model 2: Species ~ Elevation + Area
Res.Df RSS Df Sum of Sq F $\operatorname{Pr}(>F)$
$1 \quad 28 \quad 173254$
$\begin{array}{lllllll}1 & 27 & 169947 & 1 & 3307 & 0.5254 & 0.4748\end{array}$

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$P$-value is the shaped area under the under the density curve

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Multicollinearity
Another Simulated Example: Suppose the true relationship between response $Y$ and predictors ( $X_{1}, X_{2}$ ) is

$$
Y=4+0.8 X_{1}+0.6 X_{2}+\varepsilon
$$

where $\varepsilon \sim \mathrm{N}(0,1)$ and $X_{1}$ and $X_{2}$ are positively correlated with $\rho=0.9$. Let's fit the following model:

$$
Y=\beta_{0}+\beta_{1} X_{1}+\beta_{2} X_{2}+\varepsilon
$$

Call:
Imcformula
Residuals
$\begin{array}{rrrrr}\text { Min } & 10 & \text { Median } & 3 Q & \text { Max } \\ -1.63912 & -0.59978 & 0.01897 & 0.58691 & 1.74518\end{array}$
Coefficients:
Estimate Std. Error $t$ value $\operatorname{Pr}(>|t|)$
(Intercept) ${ }^{4.0154} 0.1646 \quad 24.390<2 \mathrm{e}-16$ * $\begin{array}{lrrrr}\text { x1 } & -0.1032 & 0.3426 & -0.301 & 0.766 \\ \text { X2 } & 1.7471 & 0.3654 & 4.781 & 5.48 \mathrm{e}-05 * *\end{array}$

Signif. codes: 0 '***' 0.001 '**' $0.01^{\prime * \prime} 0.05$ '.' 0.1 ', 1
Residual standard error: 0.8601 on 27 degrees of freedom Multiple R-squared: 0.8166 , Adjusted R -squared: 0.803 F-statistic: 60.12 on 2 and 27 DF, p-value: 1.135e-10

## Multicollinearity cont'd

Numerical issue $\Rightarrow$ the matrix $\boldsymbol{X}^{T} \boldsymbol{X}$ is nearly singular

- Statistical issue
- $\beta$ 's are not well estimated
- $\beta$ 's may be meaningless
- $R^{2}$ and predicted values are usually OK

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