## Lecture 4 Simple Linear Regression IV Reading: Chapter 11

STAT 8020 Statistical Methods II September 1, 2020 Simple Linear Regression IV



Analysis of Variance (ANOVA) Approach to Regression

Correlation and Coefficient of Determination

Residual Analysis: Model Diagnostics and Remedies

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### Agenda



Analysis of Variance ANOVA) Approach to Regression

Correlation and Coefficient of Determination

Residual Analysis: Nodel Diagnostics and Remedies

Analysis of Variance (ANOVA) Approach to Regression





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# ANOVA Approach to Linear Regression

#### Analysis of Variance (ANOVA) Approach to Regression

#### Partitioning Sums of Squares

• Total sums of squares in response

 $\mathsf{SST} = \sum_{i=1}^n (Y_i - \bar{Y})^2$ 

We can rewrite SST as

$$\sum_{i=1}^{n} (Y_i - \bar{Y})^2 = \sum_{i=1}^{n} (Y_i - \hat{Y}_i + \hat{Y}_i - \bar{Y})^2$$
$$= \underbrace{\sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2}_{\text{Error}} + \underbrace{\sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2}_{\text{Model}}$$

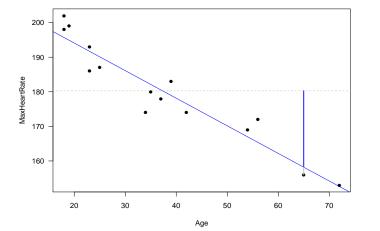




Analysis of Variance (ANOVA) Approach to Regression

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#### **Partitioning Total Sums of Squares**



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#### **Total Sum of Squares: SST**

 If we ignored the predictor X, the Y
 would be the best (linear unbiased) predictor

$$Y_i = \beta_0 + \varepsilon_i \tag{1}$$

- SST is the sum of squared deviations for this predictor (i.e., <u>Y</u>)
- The total mean square is SST/(n 1) and represents an unbiased estimate of  $\sigma^2$  under the model (1).





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#### **Regression Sum of Squares: SSR**

• SSR: 
$$\sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2$$

• Degrees of freedom is 1 due to the inclusion of the slope, i.e.,

$$Y_i = \beta_0 + \frac{\beta_1 X_i}{\lambda_i} + \varepsilon_i \tag{2}$$

• "Large" MSR = SSR/1 suggests a linear trend, because

$$E[MSE] = \sigma^{2} + \beta_{1}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}$$





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#### **Error Sum of Squares: SSE**

SSE is simply the sum of squared residuals

$$\mathsf{SSE} = \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2$$

- Degrees of freedom is n 2 (Why?)
- SSE large when |residuals| are "large" ⇒ Y<sub>i</sub>'s vary substantially around fitted regression line
- MSE = SSE/(n 2) and represents an unbiased estimate of σ<sup>2</sup> when taking X into account



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#### **ANOVA Table and F test**

Source	<b>u</b> .	SS	MS
Model	1	$SSR = \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2$	MSR = SSR/1
		$SSE = \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2$	MSE = SSE/(n-2)
Total	n-1	$SST = \sum_{i=1}^n (Y_i - \bar{Y})^2$	

- **Goal:** To test  $H_0: \beta_1 = 0$
- Test statistics  $F^* = \frac{MSR}{MSE}$
- If  $\beta_1 = 0$  then  $F^*$  should be near one  $\Rightarrow$  reject  $H_0$  when  $F^*$  "large"
- We need sampling distribution of F<sup>\*</sup> under H<sub>0</sub> ⇒ F<sub>1,n-2</sub>, where F(d<sub>1</sub>, d<sub>2</sub>) denotes a F distribution with degrees of freedom d<sub>1</sub> and d<sub>2</sub>





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#### **F Test:** $H_0: \beta_1 = 0$ **vs.** $H_a: \beta_1 \neq 0$

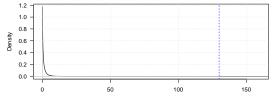
```
fit <- lm(MaxHeartRate ~ Age)
anova(fit)
....</pre>
```

A X

Analysis of Variance Table

Response: MaxHeartRate Df Sum Sq Mean Sq F value Age 1 2724.50 2724.50 130.01 Residuals 13 272.43 20.96 Pr(>F) Age 3.848e-08 \*\*\*

Null distribution of F test statistic



Test statistic

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SLR: F-Test vs. T-test

#### **ANOVA Table and F-Test**

Analysis of Variance Table

Response: MaxHeartRate Df Sum Sq Mean Sq Age 1 2724.50 2724.50 Residuals 13 272.43 20.96 F value Pr(>F) Age 130.01 3.848e-08

Parameter Estimation and T-Test

Coefficients:

 Estimate Std. Error t value Pr(>Itl)

 (Intercept) 210.04846
 2.86694
 73.27
 < 2e-16</td>

 Age
 -0.79773
 0.06996
 -11.40
 3.85e-08



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## Correlation and Coefficient of Determination

#### **Correlation and Simple Linear Regression**

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• Pearson Correlation: 
$$r = \frac{\sum_{i=1}^{n} (X_i - \bar{X})(Y_i - \bar{Y})}{\sqrt{\sum_{i=1}^{n} (X_i - \bar{X})^2 \sum_{i=1}^{n} (Y_i - \bar{Y})^2}}$$

−1 ≤ r ≤ 1 measures the strength of the linear relationship between *Y* and *X*

We can show

$$r = \hat{\beta}_1 \sqrt{\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{\sum_{i=1}^n (Y_i - \bar{Y})^2}},$$

this implies

$$\beta_1 = 0$$
 in SLR  $\Leftrightarrow \rho = 0$ 

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#### **Coefficient of Determination** *R*<sup>2</sup>

 Defined as the proportion of total variation explained by SLR

$$R^{2} = \frac{\sum_{i=1}^{n} (\hat{Y}_{i} - \bar{Y})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}} = \frac{\text{SSR}}{\text{SST}} = 1 - \frac{\text{SSE}}{\text{SST}}$$

• We can show  $r^2 = R^2$ :

$$r^{2} = \left(\hat{\beta}_{1,\text{LS}} \sqrt{\frac{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}}}\right)^{2}$$
$$= \frac{\hat{\beta}_{1,\text{LS}}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}}$$
$$= \frac{\text{SSR}}{\text{SST}}$$
$$= R^{2}$$





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Correlation and Coefficient of Determination

#### Maximum Heart Rate vs. Age: r and R<sup>2</sup>

> summary(fit)\$r.squared
[1] 0.9090967
> cor(Age, MaxHeartRate)
[1] -0.9534656

#### Interpretation:

There is a strong negative linear relationship between  ${\tt MaxHeartRate}$  and  ${\tt Age}.$  Furthermore,  $\sim 91\%$  of the variation in  ${\tt MaxHeartRate}$  can be explained by  ${\tt Age}.$ 





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Residual Analysis: Model Diagnostics and Remedies

#### Residuals

• The residuals are the differences between the observed and fitted values:

$$e_i=Y_i-\hat{Y}_i,$$

where  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$ 

- $e_i$  is NOT the error term  $\varepsilon_i = Y_i E[Y_i]$
- Residuals are very useful in assessing the appropriateness of the assumptions on ε<sub>i</sub>. Recall
  - $E[\varepsilon_i] = 0$
  - $\operatorname{Var}[\varepsilon_i] = \sigma^2$
  - $\operatorname{Cov}[\varepsilon_i, \varepsilon_j] = 0, \quad i \neq j$

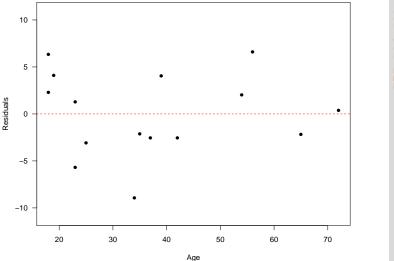




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#### Maximum Heart Rate vs. Age Residual Plot: $\varepsilon$ vs. X



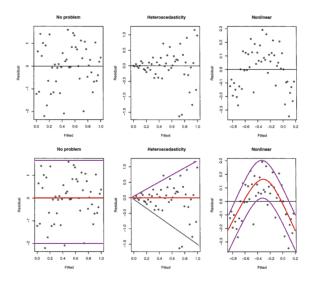
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#### **Interpreting Residual Plots**



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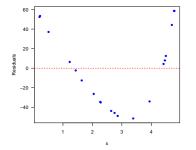
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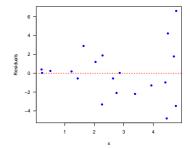
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Residual Analysis: Model Diagnostics and Remedies

**Figure:** Figure courtesy of Faraway's Linear Models with R (2005, p. 59).

#### **Model Diagnostics and Remedies**





 $\Rightarrow$  Nonlinear relationship

- $\Rightarrow$  Non-constant variance
  - Transform Y
  - Weighted least squares





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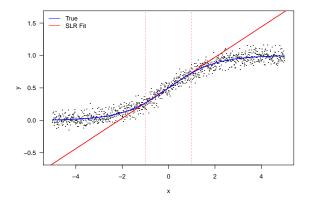
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4.20

• Transform X

Nonlinear regression



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Residual Analysis: Model Diagnostics and Remedies

Extrapolation beyond the range of the given data can lead to seriously biased estimates if the assumed relationship does not hold the region of extrapolation

#### Summary of SLR

- Model:  $Y_i = \beta_0 + \beta_1 X_i + \varepsilon_i$
- Estimation: Use the method of least squares to estimate the parameters
- Inference
  - Hypothesis Testing
  - Confidence/prediction Intervals
  - ANOVA
- Model Diagnostics and Remedies

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