Bayesian Optimization: A Brief Review

Goal: optimize $f(\mathbf{x})$ over \mathbf{x}

- $f(\mathbf{x})$ is an expensive to evaluate function.
- $f(\mathbf{x})$ is a "black-box".
- The first-order and/or second-order derivatives of f(x) is not available.

References:

- P. Frazier, "A Tutorial on Bayesian Optimization" https://arxiv.org/abs/1807.02811
- Shahriari, Bobak, et al. "Taking the human out of the loop: A review of Bayesian optimization." Proceedings of the IEEE 104.1 (2015): 148-175.

Elicit a prior distribution on the function f

while (budget is not exhausted) {

Find x that maximizes Acquisition(x, prior)

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Evaluate f(\mathbf{x}) at \mathbf{x}
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Find the posterior distribution, and update the prior distribution.

How to update knowledge, as data is obtained?

- Prior distribution: what you know about parameter β , excluding the information in the data denoted by $\pi(\beta)$.
- Likelihood: based on modeling assumptions, how [relatively] likely the data Y are if the truth is β – denoted L(Y|β)

So how to get a posterior distribution: stating what we know about β , combining the prior with the data denoted $p(\beta|Y)$. Bayes Theorem used for inference tells us to multiply:

 $p(\beta|Y) \propto L(Y|\beta)\pi(\beta)$

Essentially, Posterior \propto Likelihood \times Prior.

Given a prior distribution $\pi^{(0)}(\beta)$ for the target parameter β , and a model assumption $L(Y|\beta)$

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For t = 1, ..., N {
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obtain data Y^{(t)}
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find the posterior $p(\beta | Y^{(t)}) \propto L(Y^{(t)} | \beta) \pi^{(t-1)}(\beta)$

update $\pi^{(t)}(\beta) \leftarrow p(\beta|Y^{(t)})$

- Goal: learning parameter μ
- Prior: $\mu \sim N(\theta^{(0)}, \sigma^{(0),2})$
- Data: $Y|\mu \sim N(\mu, \lambda^2)$ where λ is known.
- Posterior: $p(\mu|Y) \propto L(Y|\mu)\pi(\mu)$ is also a normal distribution.

- Improvement-based policies: expected improvement, knowledge gradient,...
- Information-based policies: Thompson sampling

o ...

Getting into some details... with a simple example

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- Problem of Interests:

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where μ_x is the unknown true performance of alternative *x*.

 The true performance μ_x can not be directly measured, but can be estimated through observation:

$$\mathbf{y}_{\mathbf{x}}=\mu_{\mathbf{x}}+\varepsilon_{\mathbf{x}},$$

where $\varepsilon_x \sim N(0, \sigma^2)$.

Reference: A Knowledge-Gradient Policy for Sequential Information Collection P.I. Frazier, W.B. Powell & S. Dayanik. SIAM Journal on Control and Optimization, 2008.

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• Keep in mind:

 $\max_{\mathbf{X}\in\mathcal{X}}\mu_{\mathbf{X}}$

• Setup the prior belief about μ_{x}

$$\mu_{x} \sim N\left(\theta_{x}^{(0)}, (\sigma_{x}^{(0)})^{2}\right)$$

independent with each other over \mathcal{X} .

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- Assume that we collect the outputs y_{x1},..., y_{xN} are collected one by one.
- When the new observation y_{x(t)} arrives, we find the posterior distribution of μ_x given y_{x(t)}

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• In the end, find $\max_{x \in \mathcal{X}} \theta_x^{(N)}$

 Idea: choose x which provides the maximum expected "improvement" to the target problem:

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The Knowledge gradient:

$$\mathrm{KG}^{(t)}(\mathbf{X}) = \mathrm{E}[\max_{\mathbf{X}' \in \mathcal{X}} \theta_{\mathbf{X}'}^{(t+1)} - \max_{\mathbf{X}' \in \mathcal{X}} \theta_{\mathbf{X}'}^{(t)} | \mathbf{X}^{(t+1)} = \mathbf{X}],$$

where the expectation is taken with respect to the posterior predictive distribution of $Y_x^{(t+1)}$.

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• Maximize $KG^{(t)}(x)$ over \mathcal{X} to select the alternative for new experiment.

- K alternatives
- For $k=1,\ldots,K$ $\mu_k\sim N(heta_k^{(0)},\sigma_k^{(0),2})$ $Y_k|\mu_k\sim N(\mu_k,\lambda_k^2),$

where λ_k^2 is known.

- Independence between alternatives.
- Model update (if sample from the k-th alternative at step t + 1):

$$\theta_{k}^{(t+1)} = \theta_{k}^{(t)} + \frac{\sigma_{k}^{(t),2}}{\lambda_{k}^{2} + \sigma_{k}^{(t),2}} (Y_{k}^{(t+1)} - \theta_{k}^{(t)})$$
$$\sigma_{k}^{(t+1),2} = \frac{\lambda_{k}^{2} \sigma_{k}^{(t),2}}{\lambda_{k}^{2} + \sigma_{k}^{(t),2}}$$

Under the normal model with known variance, we have that

$$\begin{aligned} \mathrm{KG}^{(t)}(x) &= \mathrm{E}[\max_{x' \in \mathcal{X}} \theta_{x'}^{(t+1)} - \max_{x' \in \mathcal{X}} \theta_{x'}^{(t)} | x^{(t+1)} = x] \\ &= \tilde{\sigma}_k^{(t)} g(\xi_k^{(t)}), \end{aligned}$$

where



Expected Improvement

The expected improvement acquisition function is given by

$$\mathrm{EI}^{(t)}(\boldsymbol{x}) = \mathrm{E}\left[\max\{\mu_{\boldsymbol{x}} - \max_{j}\theta_{j}^{(t)}, \boldsymbol{0}\}\right]$$

Under the normal model,

$$\mu_{x} \sim \mathcal{N}(\theta_{x}^{(t)}, \sigma_{x}^{(t),2})$$

 $Y_{x}|\mu_{x} \sim \mathcal{N}(\mu_{x}, \lambda_{x}^{2}),$

for $k = 1, \ldots, K$. We have that,

$$\mathrm{EI}^{(t)}(x) = \sigma_x^{(t)} g\left(-\frac{|\theta_x^{(t)} - \max_j \theta_j^{(t)}|}{\sigma_x^{(t)}}\right)$$

Efficient Global Optimization, Jones et al, 1998

- Model: Gaussian process
- Acquisition function: Expected improvement

Assume

$$y(\mathbf{x}) = \mathbf{f}(\mathbf{x}_i)\boldsymbol{\beta} + \boldsymbol{\epsilon}(\mathbf{x}_i), \tag{1}$$

where $\mathbf{f}(\mathbf{x}_i) = \mathbf{f}_i$ is a pre-specified 1 × *p* regressor, β is the vector of unknown regression parameters, $\epsilon(\mathbf{x}_i)$ is a stationary Gaussian process with mean zero and covariance

$$\operatorname{cov}\left[\epsilon(\mathbf{x}_{i}), \epsilon(\mathbf{x}_{j})\right] = \sigma^{2} R(\mathbf{x}_{i}, \mathbf{x}_{j}), \text{ for } i \neq j,$$
(2)

and R is a correlation function.

- The choice of *R* determines the smoothness of $\hat{y}(\mathbf{x})$.
- One popular example:

$$R(\mathbf{x}_i, \mathbf{x}_j) = R(|\mathbf{x}_i - \mathbf{x}_j|) = \exp\left(-\sum_{k=1}^{p} \theta_k |x_{ik} - x_{jk}|^{q_k}\right), \quad (3)$$

where the subscript *k* denotes the *k*th dimension.

• Consider $R(\mathbf{h})$ for $\mathbf{h} \in \mathbb{R}^{p}$.



Comparison of different exponential power correlation functions with $\theta = 2$



Comparison of different exponential power correlation functions with q = 2







A Gaussian process with q = 1 and $\theta_k = 0.5$



Estimation of GP Parameters

- The unknown parameters involved in (1) are σ², θ = (θ₁,...,θ_d) and β.
- Given θ , the estimated σ^2 and β are

$$\hat{\boldsymbol{\beta}} = \left(\mathbf{F}^{\top}\mathbf{R}^{-1}\mathbf{F}\right)^{-1}\mathbf{F}^{\top}\mathbf{R}^{-1}\mathbf{Y}, \qquad (4)$$
$$\hat{\sigma}^{2} = \frac{(\mathbf{Y} - \mathbf{F}\hat{\boldsymbol{\beta}})^{\top}\mathbf{R}^{-1}(\mathbf{Y} - \mathbf{F}\hat{\boldsymbol{\beta}})}{n}, \qquad (5)$$

where $\mathbf{Y} = (Y_1, \dots, Y_n)$, **R** is the $n \times n$ matrix with entries $R(\mathbf{x}_i, \mathbf{x}_j)$ defined in (3) for $i, j = 1, \dots, n$ and $\mathbf{F} = [\mathbf{f}_1, \dots, \mathbf{f}_n]$.

 Given β̂ and ô², the correlation parameters θ can be estimated by maximizing the log likelihood function

$$-\frac{n}{2}\log\hat{\sigma}^2 - \frac{1}{2}\log|\mathbf{R}|.$$
 (6)

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 This model is also called Kriging, or more specific ordinary Kriging.

The BLUP predictor can be expressed by

$$\hat{\mathbf{y}}(\mathbf{x}) = \hat{\mu} + \mathbf{r}' R^{-1} (\mathbf{y} - \mathbf{1} \hat{\mu}),$$

where

•
$$\hat{\mu} = (\mathbf{1}^{\top} R^{-1} \mathbf{1})^{-1} (\mathbf{1}^{\top} R^{-1} \mathbf{y})$$

•
$$\mathbf{r} = (r(\mathbf{x}, \mathbf{x}_1), \dots, r(\mathbf{x}, \mathbf{x}_n))^\top$$

• *R* is an $n \times n$ matrix with entries $r(\mathbf{x}_i, \mathbf{x}_j)$.

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By substituting BLUP into $MSE(\hat{y}(\mathbf{x}))$, we have that

$$MSE(\hat{\boldsymbol{y}}(\mathbf{x})) = \sigma^2 \left(1 - \mathbf{r}' R^{-1} \mathbf{r} + \frac{(1 - \mathbf{1}^\top R^{-1} \mathbf{r})^2}{\mathbf{1}^\top R^{-1} \mathbf{1}} \right)$$

which is the variance of $\hat{y}(\mathbf{x})$.

An Illustration of Interpolator



Goal: $\min_{x \in \mathcal{X}} f(\mathbf{x})$, where $f(\mathbf{x})$ is a deterministic blackbox function with inputs \mathbf{x} .

Assume that the prior of $f(\mathbf{x})$ is a GP, denoted by $Y(\mathbf{x})$.

The expected improvement can be expressed by

$$\begin{split} \mathrm{EI}(\mathbf{x}) &= \mathrm{E}[\max(\mathbf{f}_{\min} - \mathbf{Y}(\mathbf{x}), \mathbf{0})] \\ &= (f_{\min} - \hat{y}(\mathbf{x})) \Phi\left(\frac{f_{\min} - \hat{y}(\mathbf{x})}{s(\mathbf{x})}\right) + s(\mathbf{x}) \phi\left(\frac{f_{\min} - \hat{y}(\mathbf{x})}{s(\mathbf{x})}\right), \end{split}$$

where $Y(\mathbf{x}) \sim N(\hat{y}(\mathbf{x}), s(\mathbf{x}))$.



Figure 11. (a) The expected improvement function when only five points have been sampled; (b) the expected improvement function after adding a point at x = 2.8. In both (a) and (b) the left scale is for the objective function and the right scale is for the expected improvement.

- We have no concave or convex property of *EI*(**x**).
- Develop a branch-and-bound algorithm to maximize *El*(x) to guaranteed optimality.

$$rac{\partial \textit{El}(\mathbf{x})}{\partial \hat{y}(\mathbf{x})} = -\Phi\left(rac{\mathit{f}_{\textit{min}} - \hat{y}(\mathbf{x})}{s(\mathbf{x})}
ight)$$

and

$$\frac{\partial EI(\mathbf{x})}{\partial s(\mathbf{x})} = \phi\left(\frac{f_{min} - \hat{y}(\mathbf{x})}{s(\mathbf{x})}\right)$$

Because of this monotonicity, to find an upper bound on *El*(**x**) over a box for **x** is suffices to find a lower bound on ŷ and an upper bound on s over the box.